COMPSCI 4ML3 Tutorial 4: Review of Probability Theory

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COMPSCI 4ML3 Tutorial 4 **Review of Probability Theory**

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Basic Elements I

- **Sample space** $Ω$: The set of all possible outcomes.
- **Event space** \mathcal{F} : The set containing all possible subsets of outcomes. i.e., A collection of possible outcomes
- **Event** A: Any element of the event space. $\forall A \in \mathcal{F}$, $A \subseteq \Omega$

For the event of rolling a dice:

 $\Omega = \{1, 2, 3, 4, 5, 6\}$

$$
\bullet\ \mathcal{F}=
$$

 $\{\{1\}, \ldots, \{6\}, \{1, 2\}, \ldots, \{5, 6\}, \{1, 2, 3\}, \ldots, \{1, 2, 3, 4, 5, 6\}\}\$

• An example of an event is $A = \{2, 3, 6\}$

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Basic Elements II

- **Probability measure** P: A funtion $P : \mathcal{F} \to \mathcal{R}$ that satisfies the following properties:
- $P(A) \geq 0, \forall A \in \mathcal{F}$
- $P(\Omega) = 1$
- For a collection of disjoint events A_i i.e., $(\forall i \neq j, \, A_i \cap A_j = \emptyset)$ we have

$$
P(\bigcup_i A_i) = \sum_i P(A_i)
$$

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Probability Measure: Properties

- If $A \subseteq B$, $P(A) \leq P(B)$
- $P(A \cup B) \leq P(A) + P(B)$, which is called Union Bound
- \bullet $P(A \cap B)$ < min($P(A), P(B)$)
- $P(A^c) = 1 P(A)$
- For disjoint events A_1, \ldots, A_k such that $\cup_{i=1}^k A_i = \Omega$

$$
\sum_{i=1}^k P(A_i) = 1,
$$

which is also called the *law of total probability*.

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Conditional Probability and Independence

• The conditional probability $P(A|B)$ is the probability of observing event A after the occurrence of B

$$
P(A|B) = \frac{P(A \cap B)}{P(B)}
$$

 \bullet Two events A and B are **independent** iff $P(A \cap B) = P(A)P(B)$. i.e, observing B does not give any information about occurrence of A and $P(A|B) = P(A)$

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Conditional Probability and Independence

Example: Probability of a person's weight being y , given that her height is x .

$$
P(\text{weight} = y | \text{height} = x)
$$

These two features are correlated.

$$
P(\text{weight} = 200\text{ lb } | \text{ height} = 190\text{ cm}) = 0.2
$$

$$
P(\text{weight} = 200\text{ lb } | \text{ height} = 140\text{ cm}) = 0.01
$$

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 $\mathcal{A} \ \equiv \ \mathcal{B} \ \ \mathcal{A} \ \equiv \ \mathcal{B}$

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Bayes' Rule

 \bullet For two events A and B

$$
P(A|B) = \frac{P(A \cap B)}{P(B)}
$$

$$
P(B|A) = \frac{P(A \cap B)}{P(A)}
$$

• This implies that

$$
P(A|B) = \frac{P(B|A)P(A)}{P(B)}
$$

$$
P(B|A) = \frac{P(A|B)P(B)}{P(A)}
$$

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Chain Rule and Law of Total Probability

• For events A_1, \ldots, A_n , chain rule states that

$$
P(A_n \cap \ldots \cap A_1) = P(A_n | A_{n-1} \cap \ldots \cap A_1) P(A_{n-1} \cap \ldots \cap A_1) =
$$

$$
P(A_1) \prod_{i=2}^n (A_i | \bigcap_{k=1}^{i-1} A_k)
$$

If B_1, \ldots, B_n are finite partition of the sample space (i.e., $\forall i\neq j, B_i \cap B_j = \emptyset$ and $\cup_{i=1}^n B_i = \Omega$), the law of total **probability** states that for an event A

$$
P(A) = \sum_{i=1}^{n} P(A \cap B_i) = \sum_{i=1}^{n} P(A|B_i)P(B_i)
$$

Random Variables

A real-valued random variable X is a mapping from sample space to real values, i.e., $X : \Omega \to \mathbb{R}$, which assigns to each element $ω ∈ Ω$ a real value $X(w)$

A random variable helps us describe some functions of observed events

- We usually denote random variables with capital letters $X(\omega)$ and simply denote it with X
- We usually use small letters for the value that a random variable may take. i.e., we write $X = x$ instead of $X(\omega) = x$

Random Variables: Example

Example: We toss coin for 20 times. What is the probability that we observe 6 heads?

- Sample space Ω can be defined as the sequences of heads and tails with length 20
- Random variable X is a function that assigns to each sequence $ω ∈ Ω$ the number of heads in that sequence. i.e., $X(\omega)$ = number of heads in ω
- We are interested in finding $P(X(\omega) = 6)$ or simply $P(X = 6)$

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Random Variables

A random variable that only takes finite number of values is called a discrete random variable

The probability that a random variable X takes value x is

$$
P(X = x) := P(\{\omega \in \Omega : X(\omega) = x\})
$$

- A random variable that can take infinite number of values is called a continuous random variable
	- **The probability that a random variable X takes values between** a and b is

$$
P(a \le X \le b) := P(\{\omega \in \Omega : a \le X(\omega) \le b\})
$$

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Cumulative Distribution Function

For a random variable X, we can define $P(X \le x)$ as a function of x:

• The Cumulative Distribution Function (CDF) is a function $F_X(x): \mathbb{R} \to [0,1]$ that is defined as

$$
F_X(x) := P(X \leq x)
$$

Properties:

\n- $$
0 \leq F_X(x) \leq 1
$$
\n- $P(a \leq X \leq b) = F_X(b) - F_X(a)$
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Cumulative Distribution Function

Example:

Probability Mass Function

For a discrete random variable, the Probability Density **Function(PMF)** $p_X(x) : \mathbb{R} \to [0,1]$ is a function that returns the probability of a random variable taking a specific value

$$
p_X(x) := P(X = x)
$$

Properties:

- 0 $0 < p_X(x) < 1$
- $\sum_{x \in \mathbb{D}} p_X(x) = 1$, where $\mathbb D$ is the set of all possible values that X can take.
- $P(X \in A) = P(\{\omega : X(\omega) \in A\}) = \sum_{x \in A} p_X(x)$

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Probability Mass Function

Example:

Probability Density Function

For a continuous random variable, we are interested in $P(x \le X \le x + \Delta x)$ when $\Delta \rightarrow 0$. If $F_X(x)$ is differentiable everywhere, the **Probability Density Function (PDF)** $f_X(x)$ is the derivative of the CDF function

$$
f_X(x) := \frac{dF_X(x)}{dx}
$$

- $P(x \le X \le x + \Delta x) \approx f_X(x) \Delta x$
- Unlike PMF, $f_X(x)$ is not the probability that the random variable X takes a value x. i.e., $f_X(x) \neq P(X = x)$. In fact, for a continuous distribution, the probability that the random variable takes a specific value is zero. i.[e,](#page-14-0) $P(X=x)=0$ $P(X=x)=0$ $P(X=x)=0$ $P(X=x)=0$ $P(X=x)=0$ $P(X=x)=0$ $P(X=x)=0$

Probability Density Function

Example:

PDF: Properties

\n- $$
f_X(x) \geq 0
$$
\n- $\int_{-\infty}^{\infty} f_X(x) = 1$
\n- $F_X(x) = \int_{-\infty}^{x} f_X(x) \, dx$
\n

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Expectation

• For a *discrete* random variable with PMF $p_X(x)$ and a function $g(x) : \mathbb{R} \to \mathbb{R}$, $g(X)$ can be considered as a random variable and the expectation or expected value of $g(X)$ is defined as

$$
\mathbb{E}[g(X)] = \sum_{x \in \mathbb{D}} g(x) p_X(x)
$$

• For a continuous random variable with PDF $f_X(x)$, the expectation or expected value of $g(X)$ is defined as

$$
\mathbb{E}[g(X)] = \int_{-\infty}^{\infty} g(x) f_X(x) dx
$$

[Expectation and Variance](#page-18-0)

Mean and Variance

Setting $g(x) = x$, the mean of a random variable X is defined as

$$
\mu = \mathbb{E}[X] = \int_{-\infty}^{\infty} x f_X(x) dx
$$

 \bullet The variance of a random variable X is a measure of how concentrated the random variable is around its mean

$$
\sigma^{2} = \text{Var} = \mathbb{E}[(X - \mathbb{E}[X])^{2}] = \mathbb{E}[X^{2} + (\mathbb{E}[X])^{2} - 2X\mathbb{E}[X]]
$$

= $\mathbb{E}[X^{2}] + (\mathbb{E}[X])^{2} - 2\mathbb{E}[X\mathbb{E}[X]] = \mathbb{E}[X^{2}] - (\mathbb{E}[X])^{2}$

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[Expectation and Variance](#page-18-0)

Mean and Variance: Example I

Example Find the mean and variance of rolling a dice with equal probability for each face

$$
\mu = \mathbb{E}[X] = \sum_{i=1}^{6} iP(X = i) = \sum_{i=1}^{6} i\frac{1}{6} = \frac{21}{6} = 3.5
$$

$$
\sigma^2 = \mathbb{E}[(X - \mu)^2] = \sum_{i=1}^{6} (i - 3.5)^2 P(X = i)
$$

$$
= \sum_{i=1}^{6} (i - 3.5)^2 \frac{1}{6} = \frac{35}{12} \approx 2.92
$$

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[Expectation and Variance](#page-18-0)

Mean and Variance: Example II

Example Find the mean and variance of a random variable with PDF $f_X(x) = 3x^2, 0 \le x \le 1$

$$
\mu = \mathbb{E}[X] = \int_{-\infty}^{\infty} x f_X(x) dx = \int_0^1 x f_X(x) dx = \int_0^1 3x^3 dx = \frac{3x^4}{4} \Big|_0^1 = \frac{3}{4}
$$

$$
\sigma^2 = \mathbb{E}[(X - \mu)^2] = \int_{-\infty}^{\infty} (x - \mu)^2 f_X(x) dx = \int_0^1 (x - \frac{3}{4})^2 3x^2 dx
$$

$$
= \int_0^1 (x - \frac{3}{4})^2 3x^2 dx = \frac{3}{16}
$$

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[Expectation and Variance](#page-18-0)

Expectation: Properties

- $\bullet \mathbb{E}[c] = c, \forall c \in \mathbb{R}$
- $\bullet \mathbb{E}[cg(X)] = c \mathbb{E}[g(X)], \forall c \in \mathbb{R}$
- $\bullet \mathbb{E}[f(X) + g(X)] = \mathbb{E}[f(X)] + \mathbb{E}[g(X)]$

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[Expectation and Variance](#page-18-0)

Variance: Properties

•
$$
Var(c) = 0, \forall c \in \mathbb{R}
$$

•
$$
Var(f(X) + c) = Var(f(X)), \forall c \in \mathbb{R}
$$

•
$$
Var(cf(X)) = c^2Var(f(X)), \forall c \in \mathbb{R}
$$

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[Discrete Random Variables](#page-24-0)

Discrete Random Variables: Bernoulli

•
$$
X \sim \text{Bernoulli}(p)
$$
, where $0 \leq p \leq 1$

$$
p_X(x) = \begin{cases} p & x = 1 \\ 1 - p & x = 0 \end{cases}
$$

\n- $$
\mathbb{E}[X] = p
$$
\n- $\text{Var}(X) = p(1 - p)$
\n

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[Discrete Random Variables](#page-24-0)

Discrete Random Variables: Binomial

•
$$
X \sim \text{Binomial}(n, p)
$$
, where $0 \le p \le 1$

$$
p_X(x) = {n \choose x} p^x (1-p)^{n-x}
$$

\n- $$
\mathbb{E}[X] = np
$$
\n- $\text{Var}(X) = np(1 - p)$
\n

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[Discrete Random Variables](#page-24-0)

Discrete Random Variables: Poisson

•
$$
X \sim \text{Possion}(\lambda)
$$
, where $\lambda > 0$

$$
p_X(x) = \frac{e^{-\lambda}\lambda^x}{x!}
$$

\n- $$
\mathbb{E}[X] = \lambda
$$
\n- $\text{Var}(X) = \lambda$
\n

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[Continuous Random Variables](#page-27-0)

Continuous Random Variables: Uniform

$$
\bullet\;\;X\sim U_{[a,b]},\;\text{where}\;a\leq b
$$

$$
f_X(x) = \begin{cases} \frac{1}{b-a} & x \in [a, b] \\ 0 & \text{otherwise} \end{cases}
$$

\n- \n
$$
\mathbb{E}[X] = \frac{b+a}{2}
$$
\n
\n- \n
$$
\text{Var}(X) = \frac{(b-a)^2}{12}
$$
\n
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[Continuous Random Variables](#page-27-0)

Continuous Random Variables: Exponential

•
$$
X \sim \text{Exponential}(\lambda)
$$
, where $\lambda > 0$

$$
f_X(x)=\lambda e^{-\lambda x}
$$

\n- $$
\mathbb{E}[X] = \frac{1}{\lambda}
$$
\n- $\text{Var}(X) = \frac{1}{\lambda^2}$
\n

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[Continuous Random Variables](#page-27-0)

Continuous Random Variables: Gaussian/Normal

•
$$
X \sim \mathcal{N}(\mu, \sigma^2)
$$

$$
p_X(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}
$$

\n- $$
\mathbb{E}[X] = \mu
$$
\n- $\text{Var}(X) = \sigma^2$
\n

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 $A \equiv \mathbf{1} \times \mathbf{1} \times \mathbf{1} \times \mathbf{1}$

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Joint Distribution

Example:

 $X = \text{image}$:

 $Y =$ label (1 if image contains cat and 0 otherwise) What is the probability of an image contains a cat?

$$
P(X = \text{image}, Y = 0) = ?
$$

$$
P(X = \text{image}, Y = 1) = ?
$$

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Joint Cumulative Distributions

It happens that we need to consider two random variables X and Y together and discuss X and Y at the same time during a random experiment.

• The joint cumulative distribution function for random variables X and Y is defined as

$$
F_{X,Y}(x,y)=P(X\leq x, Y\leq y)
$$

• The marginal CDFs can be found by

$$
F_X(x) = \lim_{y \to \infty} F_{X,Y}(x, y)
$$

$$
F_Y(y) = \lim_{x \to \infty} F_{X,Y}(x, y)
$$

[Definitions and Basics](#page-1-0) [Common Distributions](#page-24-0)
Joint Densities

[Joint CDF](#page-31-0)

Joint CDF: Properties

$$
\bullet \quad 0 \leq F_{X,Y}(x,y) \leq 1
$$

$$
\bullet \quad \lim_{x,y\to\infty} F_{X,Y}(x,y)=1
$$

$$
\bullet \quad \lim_{x,y\to -\infty} F_{X,Y}(x,y)=0
$$

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[Joint PMF](#page-33-0)

Joint Probability Mass Function

• The joint probability mass function for *discrete* random variables X and Y is defined as

$$
p_{X,Y}(x,y)=P(X=x, Y=y)
$$

• The marginal PMFs can be found by

$$
p_X(x) = \sum_{y \in \mathbb{D}_y} p_{X,Y}(x,y)
$$

$$
p_Y(y) = \sum_{x \in \mathbb{D}_y} p_{X,Y}(x,y)
$$

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[Definitions and Basics](#page-1-0) [Common Distributions](#page-24-0)
Joint Densities

[Joint PMF](#page-33-0)

Joint PMF: Properties

$$
\bullet \quad 0 \leq p_{X,Y}(x,y) \leq 1
$$

$$
\bullet \sum_{x \in \mathbb{D}_x, y \in \mathbb{D}_y} p_{X,Y}(x,y) = 1
$$

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Joint Probability Density Function

If the joint CDF is differentiable everywhere in x and y, the joint probability density function for *continuous* random variables X and Y is defined as

$$
f_{X,Y}(x,y) = \frac{\partial^2 F_{X,Y}(x,y)}{\partial x \partial y}
$$

• The marginal PDFs can be found by

$$
f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dy
$$

$$
f_Y(y) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dx
$$

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[Definitions and Basics](#page-1-0) [Common Distributions](#page-24-0)
Joint Densities

[Joint PDF](#page-35-0)

Joint PDF: Properties

$$
\bullet \quad \int_{-\infty}^{\infty}\int_{-\infty}^{\infty}f_{X,Y}(x,y)dxdy=1
$$

$$
\bullet \quad \int\int_A f_{X,Y}(x,y)dxdy = P((X,Y) \in A)
$$

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[Conditional PMF and PDF](#page-37-0)

Conditional Distributions

• Conditional PMF refers to the probability distribution over X when we know that Y has taken a certain value (if $p_Y(y) \neq 0$)

$$
p_{X|Y}(x|y) = P(X = x|Y = y) = \frac{P(X = x, Y = y)}{P(Y = y)} = \frac{p_{X,Y}(x,y)}{p_Y(y)}
$$

• Conditional PDF is defined as (if $f_Y(y) \neq 0$)

$$
f_{X|Y}(x|y) = \frac{f_{X,Y}(x,y)}{f_Y(y)}
$$

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 $\mathcal{A} \ \equiv \ \mathcal{B} \ \ \mathcal{A} \ \equiv \ \mathcal{B}$

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[Independence](#page-38-0)

Independent Random Variables

Two random variables X and Y are independent iff

$$
F_{X,Y}(x|y) = F_X(x)F_Y(y), \forall x, y
$$

 \bullet If two discrete random variables X and Y are independent

$$
p_{X,Y}(x,y) = p_X(x)p_Y(y), \forall x, y
$$

\n
$$
p_{X|Y}(x|y) = p_X(x), \forall x, y \text{ such that } p_Y(y) \neq 0
$$

 \bullet If two continuous random variables X and Y are independent

$$
f_{X,Y}(x,y) = f_X(x) f_Y(y), \forall x, y
$$

$$
f_{X|Y}(x|y) = f_X(x), \forall x, y \text{ such that } f_Y(y) \neq 0
$$

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[Bayes' Rule](#page-39-0)

Bayes' Rule for Joint Probability Distribution

• For two *discrete* random variables X and Y

$$
p_{X|Y}(x|y) = \frac{p_{Y|X}(y|x)p_X(x)}{p_Y(y)}
$$

• For two continuous random variables X and Y

$$
f_{X|Y}(x|y) = \frac{f_{Y|X}(y|x)f_X(x)}{f_Y(y)}
$$

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[Joint Densities](#page-30-0) [Expectation and Covariance](#page-40-0)

Expectation of Joint Distributions

 \bullet For discrete random variables X and Y with joint PMF $p_{X,Y}(x,y)$ and a function $g(x,y):\mathbb{R}^2\rightarrow\mathbb{R}$, $g(X,Y)$ can be considered as a random variable and the expectation or expected value of $g(X, Y)$ is defined as

$$
\mathbb{E}[g(X,Y)] = \sum_{x \in \mathbb{D}_x} \sum_{y \in \mathbb{D}_y} g(x,y) p_{X,Y}(x,y)
$$

 \bullet For continuous random variables X and Y with joint PDF $f_{X,Y}(x, y)$, the expectation or expected value of $g(X, Y)$ is defined as

$$
\mathbb{E}[g(X,Y)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x,y) f_{X,Y}(x,y) dxdy
$$

[Expectation and Covariance](#page-40-0)

Covariance of Joint Distributions

 \bullet The covariance of two random variables X and Y is defined as

$$
Cov(X, Y) = \mathbb{E}[(X - \mathbb{E}[X])(Y - \mathbb{E}[Y])]
$$

= $\mathbb{E}[XY - Y\mathbb{E}[X] - X\mathbb{E}[Y] + \mathbb{E}[X]\mathbb{E}[Y]]$
= $\mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y] - \mathbb{E}[X]\mathbb{E}[Y] + \mathbb{E}[X]\mathbb{E}[Y]$
= $\mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y]$

If $Cov(X, Y) = 0$, two random real-valued variables are called uncorrelated.

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[Expectation and Covariance](#page-40-0)

Expectation and Covariance: Properties

- $\bullet \quad \mathbb{E}[f(X, Y) + g(X, Y)] = \mathbb{E}[f(X, Y)] + \mathbb{E}[g(X, Y)]$
- \bullet Var $(X + Y) = \text{Var}(X) + \text{Var}(Y) + 2\text{Cov}(X, Y)$
- If X and Y are independent, $Cov(X, Y) = 0$ \bullet
- **If** X and Y are independent, $\mathbb{E}[f(X)g(Y)] = \mathbb{E}[f(X)\mathbb{E}[g(Y)]$

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[Multivariate distributions](#page-43-0)

Generalized Joint Distribution for n Variables

For *n* random variables X_1, \ldots, X_n

Q Joint CDF is defined as

$$
F_{X_1,\ldots,X_n}(x_1,\ldots,x_n)=P(X_1\leq x_1,\ldots,X_n\leq x_n)
$$

• For *discrete* random variables joint PMF is defined as

$$
p_{X_1,...,X_n}(x_1,...,x_n) = P(X_1 = x_1,...,X_n = x_n)
$$

• For continuous random variables joint PDF is defined as

$$
f_{X_1,\ldots,X_n}(x_1,\ldots,x_n)=\frac{\partial^n F_{X_1,\ldots,X_n}(x_1,\ldots,x_n)}{\partial x_1\ldots\partial x_n}
$$

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[Multivariate distributions](#page-43-0)

Joint Distribution for n Variables

For *n* randaom variable X_1, \ldots, X_n

• Marginal PDFs can be derived by

$$
f_{X_1}(x_1)=\int_{-\infty}^{\infty}\ldots\int_{-\infty}^{\infty}f_{X_1,\ldots,X_n}(x_1,\ldots,x_n)dx_2\ldots dx_n
$$

 \bullet

$$
P((X_1,\ldots,X_n)\in A)=\int\ldots\int_A f_{X_1,\ldots,X_n}(x_1,\ldots,x_n)dx_1\ldots dx_n
$$

 \bullet X_1, \ldots, X_n are mutually independent iff

$$
f_{X_1,\ldots,X_n}(x_1,\ldots,x_n)=\prod_{i=1}^n f_{X_i}(x_i)
$$

[Random Vectors](#page-45-0)

Random Vectors

When dealing with *n* random variables, we can consider them as a random vector $X=[X_1,\ldots,X_n]^T$

• For the random vector X , the expectation is in the form of a vector. For a function $g: \mathbb{R}^n \to \mathbb{R}^m$

$$
\mathbb{E}[g(X)] = \begin{bmatrix} \mathbb{E}[g_1(X)] \\ \vdots \\ \mathbb{E}[g_m(X)] \end{bmatrix}
$$

• The mean vector is $\mu = \mathbb{E}[X] = [\mathbb{E}[X_1], \dots, \mathbb{E}[X_n]]^T$

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Covariance Matrix

For a random vector $X \in \mathbb{R}^n$, its covariance matrix $\Sigma \in \mathbb{R}^{n \times n}$ is a symmetric positive semidefinite matrix, where $\Sigma_{i,j} = \mathsf{Cov}(X_i, X_j)$

$$
\Sigma = \begin{bmatrix}\n\text{Cov}(X_1, X_1) & \text{Cov}(X_1, X_2) & \dots & \text{Cov}(X_1, X_n) \\
\text{Cov}(X_2, X_1) & \text{Cov}(X_2, X_2) & \dots & \text{Cov}(X_2, X_n) \\
\vdots & \vdots & \ddots & \vdots \\
\text{Cov}(X_n, X_1) & \text{Cov}(X_n, X_2) & \dots & \text{Cov}(X_n, X_n)\n\end{bmatrix}
$$
\n
$$
= \mathbb{E}[(X - \mathbb{E}[X])(X - \mathbb{E}[X])^T] = \mathbb{E}[XX^T] - \mathbb{E}[X]\mathbb{E}[X]^T
$$

 $\mathsf{Note}\ \mathsf{Cov}(X_i,X_i)=\mathsf{Var}(X_i)$

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Multivariate Gaussian Distribution

A multivariate Gaussian random variable $X \sim \mathcal{N}(\mu, \Sigma)$ can be defined as

$$
f_{X_1,...,X_n}(x_1,...,x_n)=\frac{1}{\sqrt{(2\pi)^n}|\Sigma|^{1/2}}\exp(-\frac{1}{2}(x-\mu)^T\Sigma^{-1}(x-\mu))
$$

If variables X_1, \ldots, X_n are uncorrelated, the covariance matrix Σ will become a diagonal matrix with variances of individual variables in its main diagonal. In this case,

$$
f_{X_1,...,X_n}(x_1,...,x_n) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi}\sigma_i} \exp(-\frac{(x_i - \mu_i)^2}{2\sigma_i^2})
$$

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